## CHAPTER FOUR

### METHODLOGY

# 4.1 Introduction

In this chapter, the core econometric methods that will be used in estimating the effect of fiscal policy on the real exchange rates and trade balance will be discussed. These methods include the unit root tests, which is used to examine if a variable is stationary or nonstationary. The cointegration technique on the other hand will be applied to estimate the long – run relationship between fiscal policy, real exchange rates and the trade balance. This technique would be using the Engle – Granger approach.

## 4.1.1 Unit root tests

A time series data can be thought of as being generated by a stochastic or random process. A stochastic process is said to be stationary if its mean and variance are constant overtime and the value of covariance between two time periods depends only on the distance or lag between the two time periods and not on the actual time at which the covariance is computed. If we assume Yt to be a stochastic time series it will therefore have the following time series properties.

Mean:	$E(Y_t) = \mu$		(4.1	)
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Variance:  $var(Y_{t}) = E(Y_{t} - \mu)^{2} = \sigma_{2}$  (4.2) Covariance:  $\gamma_{k} = E((Y_{t} - \mu)(Y_{t} + k - \mu))$  (4.3)

Where  $\gamma_k$ , the covariance (or autocovariance) at lag k, is the covariance a between the values of Y<sub>1</sub> and Y<sub>1</sub> + k, that is between two Y values k periods apart. If k = 0, we obtain  $\gamma_0$ , which is simply the variance of Y(=  $\sigma^2$ ), if k = 1,  $\gamma_1$  is the covariance between two adjacent values of Y.

If we shift the origin of Y from Y<sub>t</sub> to Y<sub>t</sub> + m, now if Y<sub>t</sub> is said to be stationary, the mean, the variance and autocovariances of Y<sub>t</sub> + m must be the same as those of Y<sub>t</sub>. In short if a time series is stationary, its mean variance and autocovariances (at various lags) remain the same no matter at what time it is measured. If a time series is not stationary, it is called a nonstationary time series. At times this nonstationary could be due to a shift in the mean. In order to see if a time series is stationary or nonstationary, a unit root test is applied.

Using a simple model as below:

 $Y_t = Y_{t-1} + \mu_t$  (4.4)

Where  $\mu_i$  is the stochastic term that follows the classical assumptions, namely it has zero mean, constant variance  $\sigma^2$ , and is non autocorrelated. Such an error term is also known as a white noise error term. Equation (4.4) can also be considered as an first order or autoregressive regression in that the value of Y at time t is regressed against the values of Y at time (t-1). If the coefficient of  $Y_{t-1}$  is in fact equal to 1, therefore a unit root problem exist, or in other words a nonstationary situation exist or in otherwords the null hyphothesis. Therefore the following regression is run:

$$Y_t = pY_{t-1} + \mu_t$$
 (4.5)

If in equation (4.5), the regression that is run, obtains a value of p = 1, therefore it can be considered that variable Yt has a unit root. In time series econometrics, a time series that has a unit root is known as a random walk time series, and a random walk is an example of a nonstationary time series. Equation (4.5) is often expressed in an alternative form as:

Where  $\delta = (p - 1)$  and where  $\blacktriangle$  is the first difference operator or in otherwords

 $\blacktriangle Yt = (Yt - Yt - 1).$ 

Making use of this definition, the null hyphothesis is  $\delta = 0$  while the alternative hyphothesis is  $\delta \# 0$ . If it is in fact 0, therefore equation (4.6) can be written as:

 $\mathbf{A} Y_t = (Y_t - Y_{t-1}) = \mu_t \tag{4.7}$ 

Equation (4.7) says that the first difference of a random walk time series (=  $\mu$ ) are a stationary time series because by assumption  $\mu$ t is purely random. If a time series is differenced once and the differenced series is stationary, therefore the original (random walk) series is integrated of order 1, denoted by I(1). Similarly if the original series has to be differenced twice (i.e. take first difference of the first difference), before it becomes stationary, the original series is integrated of order 2 or I(2). In general, if a time series has to be differenced 'd' times, it is , integrated of order 'd' or I(d).

To find out if a time series Yt is nonstationary, firstly regression of equation (4.5) must be run, in order to find out if 'p' is statistically equal to 1 or, equivalently estimate equation (4.6) if  $\delta = 0$ . Using the Dickey – Fuller test (tau test) the null hyphotesis, of p = 1 and  $\delta = 0$  is rejected if the variable is non-stationary.

In its simplest form, a regression equation (4.5) is estimated, divide the estimated "p" coefficient by its standard error to compute the Dickey -Fuller tau statistics and later refer it to the Dickey - Fuller tables to see if the null hyphothesis p = 1 is rejected. If the computed absolute value of the tau statistics (i.e.[T]) exceeds the critical value in the Dickey -Fuller critical values, then we reject the hyphothesis that the given time series is non-stationary. If on the other hand, if it is less than the critical value, the time series is non-stationary and we accept the alternative hyphothesis that the given time series is non-stationary. The Dickey-Fuller test is applied to regressions run in the following form:

 $AY_t = \delta Y_{t-1} + \mu t$  (4.8)

 $AY_t = \beta 1 + \delta Y_{t-1} + \mu t - (4.9)$ 

$$\Delta Y_t = \beta 1 + \beta 2 t + \delta Y_{t-1} + \mu t \quad (4.10)$$

Equation (4.8) excludes the constant and time trend variable, while equation (4.9) excludes the time trend variable only, on the other hand equation (4.10) includes both the time trend and and constant variable. In each case the null hyphothesis is that  $\delta = 0$ , that is there is a unit root. However, if the error term ( $\mu$ ) is autocorrelated, therefore equation  $_{\chi}(4.10)$  would have to be modified as follows:

$$Y_t = \beta 1 + \beta 2 \text{Trend} + \delta Y_{t-1} + \alpha_i \sum A Y_{t-i} + \varepsilon_t$$
(4.11)

Where  $\mathbf{A}Yt-1 = (Yt-1 - Yt-2)$  and  $\mathbf{A}Yt-2 = (Yt-2 - Yt - 3)$  and so forth, that is one uses lagged difference terms. The number of lagged difference terms to include is often determined empirically, the idea being to include enough terms so that the error term in equation (4.11) is serially independent. The null hyphothesis is still that  $\delta = 0$  or  $\mathbf{p} = 1$ that is a unit root exists in Y (i.e. Y is nonstationary ). When the Dickey – Fuller test is applied to models like equation (4.11) it is called as Augmented Dickey – Fuller (ADF) test.

## 4.2 Testing for Cointegration: The Engle - Granger method.

Suppose that two variables such as  $Y_t$  and  $Z_t$  are believed to be integrated of order one and if we want to determine there exists an equilibirium relationship between the two. Engle and Granger (1987) propose a straightforward test whether the two I(1) variables are or  $\dot{f}$  order CI(1, 1). Step 1: Pretest the variables for their order of integration. By definition, cointegration necessitates that the variables be integrated of the same order. Thus the first step in the analysis is to pretest each variable to determine its order of integration. The Dickey – Fuller tests can be used to infer the number of unit roots (if any) in each of the availables. If both variables are stationary, it is not necessary to proceed since standard time series methods apply to stationary variables. If the variables are integrated of different orders, it is possible to conclude that they are not cointegrated.

Step 2 : Estimate the long – run equilibirium relationship. If the results of step 1 indicate that both Yt and Zt are I(1), the next step is to estimate the long – run equilibirium relationship in the form:

$$Y_t = \beta_0 + \beta_1 Z_t + \epsilon_t$$
 (4.12)

If the variables are cointegrated, an OLS regression would provide a super consistent estimator of the cointegrating parameters of  $\beta 0$  and  $\beta 1$ . According to Stock (1987), OLS models estimates of  $\beta 0$  and  $\beta 1$  would converge if the variables are stationary. In order to determine if the variables are actually cointegrated denote the residual sequence by  $|\epsilon_t|$ . Thus  $|\epsilon_t|$ , is the series of the estimated residuals of the long – run relationship. If these deviations from long – run equilibirium are found to be stationary, the Yt and Zt sequences are cointegrated of order (1, 1). It would be convenient if a Dickey – Fuller test is performed on these residuals to determine their order of integration.

Considering the autoregression of the residuals;

 $A \in t = \alpha 1 \in t - 1 + \psi t$  (4.13)

Since the |et | value is a residual from a regression equation, there is no need to include an intercept term. The parameter of interest in the equation above is  $\alpha 1$ . If we cannot reject the null hyphothesis  $\alpha 1 = 0$ , therefore we can conclude that the residual series contains a unit root. Hence we conclude that the Yt and Zt variables are not cointegrated. The more precise wording is not appropriate because of a triple negative, but to be technically correct, if it is not possible to reject the null hyphothesis  $|\alpha_1| = 0$ , we cannot reject the hyphothesis that the variables are not co-integrated. Instead, the rejection of the null hyphothesis implies that the residual sequence is stationary. Given that both Yt and Zt were found to be I(1) and the residuals are stationary, we can conclude that the series are cointegrated of order (1, 1). If the residuals from equation (4.13) do not appear to be white noise, an augmented Dickey - Fuller test can be used instead of (4.13). Suppose that diagnostic checks indicate that wt sequence of (4.13) exhibits serial correlation. Instead of using the results from (4.13), estimate the autoregression (4.14) in order to conclude that the residual sequence is stationary and  $Y_t$  and  $Z_t$  are cointegrated in the order (1, 1):

$$A \in t = \alpha 1 \in t - 1 + \sum \alpha i \cdot 1 A \in t - 1 + \psi t$$
 ------ (4.14)

Step 3: Estimate the error-correction model. If the variables are cointegrated (i.e. if the null hyphothesis of no cointegration is rejected), the residuals from the equilibrium regression can be used to estimate the error-correction model. The value of the residuals  $\epsilon_{t-1}$  estimates the deviation from long-run equilibrim in period (t-1). Hence it is possible to use the saved residuals  $(\epsilon_{t-1})$  obtained in Step 2 as an instrument to estimate the following error-correction models.

- Lt = random disturbance
- Kt = random disturbance
- t-i = lagged quarters

The speed of adjustment coefficients  $\alpha_2$  and  $\beta_2$  are of particular interest in that they have important implications for the dynamics of the system. From equation (4.16), it is clear that for any given value of  $\epsilon_{i+1}$ , a large value of  $\beta_2$  is associated with a large value of  $\Delta Z_i$ . If  $\beta_2$ is zero, the change in  $Z_i$  does not at all respond to the deviation from long-run equilibrium in (t-1). If  $\beta_2$  is zero and  $\beta_3 = 0$ , then it can be said that  $\Delta Y_t$  does not Granger cause  $\Delta Z_t$ . One or both of these coefficients should be significantly different from zero if the variables are cointegrated. After all if both  $\alpha_2$  and  $\beta_2$  are zero, there is no errorcorrection and equation (4.15) and (4.16) comprise nothing more than a Vector Autoregressive in first differences. Moreover, the absolute values of these speed of adjustment coefficients must not be too large. The point estimates should imply that  $\Delta Y_t$  and  $\Delta Z_t$  converge to the long-run equilibrium relationship.